



Let a domain V in an orebody be the union $V = \bigcup v_i$ of N small units v_i , all disjoint and equal to a same volume v (up to a translation). It seems natural to characterize the dispersion of grades $z(v_i)$ of these units v_i inside the domain V by the parameter

$$s^{2}(v \mid V) = \frac{1}{N} \sum \left[z(v_{i}) - z(V) \right]^{2}$$

Unfortunately, grades $z(v_i)$ and z(V) generally are not known, and parameter $s^2(v \mid V)$ cannot be estimated without using a model. Many models are used in geostatistics. In the most common one, grades z(x) (defined on a point support) are interpreted as a realization of an intrinsic random function Z(x) (of order 0). Within the framework of this model, the parameter $s^2(v \mid V)$ may be interpreted as a realization of the corresponding random variable

$$S^{2}(v \mid V) = \frac{1}{N} \sum \left[Z(v_{i}) - Z(V) \right]^{2}$$

and the expectation of this random variable is easy to calculate with the help of the variogram $\gamma(h)$, i.e.

$$\sigma^{2}(v \mid V) = E[S^{2}(v \mid V)] = \overline{\gamma}(V, V) - \overline{\gamma}(v, v)$$
 (1)

In this expression, $\gamma(V, V)$ represents the average value of $\gamma(x - y)$ when the points x and y are distributed uniformly inside V. This new parameter $\sigma^2(v \mid V)$ is called "dispersion variance" (of v inside V). Its numerical value depends on the geometry (v and V) and also on the choice of the variogram $\gamma(h)$. When available data allow a correct choice of the variogram, formula (1) generally gives excellent results, i.e., good agreement with the average value of $s^2(v \mid V)$ whenever it may be calculated for different domains with same size V inside the same orebody.

All of this is elementary and fully explained in French: Matheron (1962, p. 57-61 and 72-73; 1965, p. 135-140); in Russian: Matheron (1968, p. 94-

100); in English: Matheron (1970, p. 66-69), David (1977, p. 97-98), Journel and Huijbregts (1978, p. 61-68), and now in many other languages.

In the particular case of a logarithmic variogram (the "de Wijsian scheme" of the sixties) and if volumes v and V are similar (i.e., have the same shape, but not the same size!), the general formula (1) becomes

$$\sigma^{2}(v \mid V) = \alpha \log (V/v)$$
 (2)

This can be found in Matheron (1962, p. 76; 1965, p. 242; 1968, p. 98) and in David (1977, p. 180).

Concerning the original de Wijs's model, it dates back to a pregeostatistical era. It has no random function in it. Starting from a given unit V with grade Z(V) = m, and dividing it (in mind!) into two equal parts V/2, de Wijs obtains two grades, say m(1+d) and m(1-d) ($0 \le d \le 1$). By iterating this process k times, and assuming (as a first approximation!) that d remains constant, this results in 2^k small units v; among them

$$\binom{k}{n}$$
 have grade $(m(1-d)^{k-n}(1+d)^n$

 $(0 \le n \le k)$. Thus, if we choose at random one of these units v among the 2^k existing ones, its grade Z(v) is a random variable with a logbinomial distribution; in fact, we have

$$\log Z(v) = \log m + k \log (1 - d) + n \log [(1 + d)/(1 - d)]$$

and n is binomial. The variance of n is k/4, and thus

$$\operatorname{Var}\left[\log Z(v)\right] = \frac{k}{4}\log\left(\frac{1+d}{1-d}\right)^{2}$$

Moreover, V being divided 2^k parts v, we have $V/v = 2^k$, i.e.

$$k = \log (V/v)/\log 2$$

It follows

Var
$$\left[\log Z(v)\right] = \alpha \log \left(V/v\right)$$
 with
$$\alpha = \frac{1}{4 \log 2} \left[\log \left(\frac{1+d}{1-d}\right)\right]^2 \tag{3}$$

Equations (2) and (3) have the same form. Nevertheless, the first gives the variance of the grade Z(v), whereas the second concerns the variance of its logarithm.

In de Wijs's logbinomial model, the variance of Z(v) is

$$\operatorname{Var}\left[Z(v)\right] = m^{2}\left[\left(V/v\right)^{\beta} - 1\right] \quad \text{with}$$

$$\beta = \log\left(1 + d^{2}\right)/\log 2 \tag{4}$$

Letter to the Editor 457

From the geostatistical point of view of formula (1), formula (4) corresponds to the choice of a variogram of the form

$$\gamma(h) = A \mid h \mid^{r} \tag{5}$$

with $v = n\beta$ if the space has *n* dimensions. The necessary condition 0 < v < 2 is satisfied automatically for n = 1 or 2, because we have $0 < \beta < 1$, but not for n = 3.

This connection between de Wijs's original model and geostatistics is explained in Matheron (1962, p. 308-311; 1968, p. 322-325). The interpretation of formula (4) with the help of a variogram of the form (5) is given in Matheron (1965, p. 139; 1968, p. 98).

In spite of their particular character, the "de Wijsian" equation (3) and (4) remain relatively interesting today because they express what I called a "similarity principle" (Matheron, 1962, p. 76). In more up-to-date terminology, one would speak of "fractal processes." In fact, "self-similar processes" is better because the term "fractal" is now becoming very misleading. But, whatever their names, these processes are old acquaintances for geostatisticians.

G. Matheron

Centre de Geostatistique

Ecole des Mines de Paris

35, Rue Saint-Honore 35

77305 Fontainebleau, France

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